ECO 519. Handout on Quantile Regression

1 Setup

The classic reference is Koenker and Basset (1978). We start by noting the result that suppose a random variable z_i has distribution given by $F_z(\cdot)$. Take $\tau \in (0, 1)$, then the solution to the problem

Minimize
$$E\left[\tau \cdot |z_i - b| \mathbf{1}\{z_i - b \ge 0\} + (1 - \tau) \cdot |z_i - b| \mathbf{1}\{z_i - b < 0\}\right]$$

is given by $b = F_z^{-1}(\tau)$, or the τ^{th} quantile of z_i . We can show this immediately if we assume that z_i has a continuous distribution:

$$E\Big[\tau \cdot |z_i - b| \mathbb{1}\{z_i - b \ge 0\} + (1 - \tau) \cdot |z_i - b| \mathbb{1}\{z_i - b < 0\}\Big] = \tau \int_b^\infty (z - b) f_z(z) dz - (1 - \tau) \int_{-\infty}^b (z - b) f_z(z) dz$$

and taking first order conditions with respect to b yields (via Leibniz' Rule) $F_z(b) = \tau$, or $b = F_z^{-1}(\tau)$. Now consider the linear model

$$y_i = \beta'_0 x_i + u_i,$$

where $u_i \sim F_u(\cdot)$ is continuously distributed and independent of x_i . The τ^{th} -quantile of u_i is given by $F_u^{-1}(\tau)$ independently of x_i . Suppose from the onset that x_i includes a constant (i.e, an intercept). Consider now the problem

$$\underset{\beta}{\text{Minimize } E\left[\tau \cdot |y_i - x_i'\beta| \mathbb{1}\{y_i - x_i'\beta \ge 0\} + (1 - \tau) \cdot |y_i - x_i'\beta| \mathbb{1}\{y_i - x_i'\beta < 0\} |x_i|\right].$$

This is simply

$$\underset{\beta}{\min} E\Big[\tau \cdot |u_i - x_i'(\beta - \beta_0)| \mathbf{1}\{u_i - x_i'(\beta - \beta_0) \ge 0\} + (1 - \tau) \cdot |u_i - x_i'(\beta - \beta_0)| \mathbf{1}\{u_i - x_i'(\beta - \beta_0) < 0\} |x_i].$$

As we saw above, the solution to this problem must be the conditional τ^{th} -quantile of u_i given x_i . That is, for all x_i we must have

$$x'_i(\beta - \beta_0) = F_u^{-1}(\tau) \implies x'_i\beta = x'_i\beta_0 + F_u^{-1}(\tau),$$

which (given the fact that x_i includes an intercept) is solved by using $\beta = \beta_0 + \iota F_u^{-1}(\tau)$, where $\iota = (1, 0, ..., 0)$. We denote this shortly by $\beta(\tau)$ and assume that it is the unique solution (think about the usual full-rank assumption). Note that if the τ^{th} quantile of u_i is zero, then $\beta(\tau) = \beta_0$ and most importantly, notice that $\beta(\tau)$ and β_0 differ only in the intercept. Notice that all the action takes place in the intercept.

2 Estimation

Using the analogy principle, we can estimate $\beta(\tau)$ by solving

$$\min_{\beta} \frac{1}{N} \sum_{i=1}^{N} \Big(\tau \cdot |y_i - x_i'\beta| \mathbb{1}\{y_i - x_i'\beta \ge 0\} + (1-\tau) \cdot |y_i - x_i'\beta| \mathbb{1}\{y_i - x_i'\beta < 0\} \Big),$$

which can be simplified to

$$\operatorname{Min}_{\beta} \frac{1}{N} \sum_{i=1}^{N} |y_i - x'_i \beta| \cdot [\tau + (1 - 2\tau) \mathbf{1} \{y_i - x'_i \beta < 0\}]$$

After minor simplification, the left-partial derivative with respect to β_{ℓ} is easily given by

$$S_{N,\ell}(\beta) \equiv \frac{1}{N} \sum_{i=1}^{N} x_{i,\ell} (1 \{ y_i < x'_i \beta \} - \tau).$$

Now remember the key argument we used when discussing Powell's Censored LAD. Since $\hat{\beta}$ is minimizing the objective function, it must be case that for any $\delta > 0$:

$$S_{N,\ell}(\widehat{\beta}_{\ell}-\delta,\widehat{\beta}_{-\ell}) \leq S_{N,\ell}(\widehat{\beta}) \leq S_{N,\ell}(\widehat{\beta}_{\ell}+\delta,\widehat{\beta}_{-\ell}) \implies \left|S_{N,\ell}(\widehat{\beta})\right| \leq \left|S_{N,\ell}(\widehat{\beta}_{\ell}+\delta,\widehat{\beta}_{-\ell}) - S_{N,\ell}(\widehat{\beta}_{\ell}-\delta,\widehat{\beta}_{-\ell})\right|$$

and this holds for all $\ell = q, \ldots, K \equiv \dim(\beta)$. This implies that

$$\begin{split} \left| S_{N,\ell}(\widehat{\beta}) \right| &\leq \frac{1}{N} \sum_{i=1}^{N} \left| x_{i,\ell} \left(\mathbf{1}\{y_i < x'_i \widehat{\beta} + x_{i,\ell}\} - \mathbf{1}\{y_i < x'_i \widehat{\beta} - x_{i,\ell}\} \right) \right| \\ &= \frac{1}{N} \sum_{i=1}^{N} \left| x_{i,\ell} \right| \cdot \mathbf{1}\{x'_i \widehat{\beta} - |x_{i,\ell}| \delta < y_i < x'_i \widehat{\beta} + |x_{i,\ell}| \delta \} \end{split}$$

Since this holds for any $\delta > 0$, letting $\delta \to 0$ we get

$$\sqrt{N} |S_{N,\ell}(\widehat{\beta})| \le \max_{i=1,\dots,N} |x_{i,\ell}| \frac{1}{\sqrt{N}} \sum_{i=1}^{N} \mathbb{1}\{y_i = x_i'\widehat{\beta}\} = \max_{i=1,\dots,N} |x_{i,\ell}| \frac{1}{\sqrt{N}} \sum_{i=1}^{N} \mathbb{1}\{u_i = x_i'(\widehat{\beta} - \beta_0)\}$$

If we assume that $||x_i||$ is well-behaved in the sense for example that there exists a random variable z_i such that $||x_i|| \le z_i$ w.p.1, and $E|z_i| < \infty$, then $S_N(\beta)$ is easily uniformly bounded

by $\frac{1}{N} \sum_{i=1}^{N} z_i$. This yields uniform convergence of $S_N(\beta)$ to $E\left[x_i\left(F_u(x'_i(\beta-\beta_0)-\tau)\right)\right]$, which is uniquely minimized at $\beta(\tau)$. The continuously-distributed nature of u_i and the fact that $\widehat{\beta} \xrightarrow{p} \beta(\tau)$ yields $\frac{1}{\sqrt{N}} \sum_{i=1}^{N} \mathbb{1}\{u_i = x'_i(\widehat{\beta} - \beta_0)\} = o_p(1)$. Therefore

$$\underbrace{\max_{i=1,\dots,N} |x_{i,\ell}|}_{=O_p(1)} \underbrace{\frac{1}{\sqrt{N}} \sum_{i=1}^N \mathbb{1}\{u_i = x'_i(\widehat{\beta} - \beta_0)\}}_{=o_p(1)}}_{=o_p(1)} = o_p(1)$$

and therefore $\sqrt{N}S_{N,\ell}(\widehat{\beta}) = o_p(1)$. To find the asymptotic distribution of $\sqrt{N}(\widehat{\beta} - \beta(\tau))$, the best and fastest thing to do is to analyze the empirical process:

$$\nu_{N}(\beta) = \frac{1}{\sqrt{N}} \sum_{i=1}^{N} \left[x_{i} \left(\mathbf{1} \{ y_{i} < x_{i}'\beta \} - \tau \right) - E \left[x_{i} \left(\mathbf{1} \{ y_{i} < x_{i}'\beta \} - \tau \right) \right] \right]$$

This is easily a vector-valued empirical process spanned by Type-I functions (Andrews, pp.2270). In addition, the empirical process $\nu_N(\beta)$ has envelope z_i , so if we assume $E[z_i^{2+\delta}] < \infty$ for some $\delta > 0$, then all the assumptions in Theorem 1 in Andrews. Therefore $\nu_N(\beta)$ is stochastically equicontinuous. Therefore, the asymptotic distribution of $\sqrt{N}(\hat{\beta} - \beta(\tau))$ can be immediately derived as in Section 3.2 in Andrews.